

Chapter 2 Equations and Algebra

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Key-words

- Linear model – method of least squares, normal equations
- Inverse matrix
- Rank
- Non-singularity
- Determinant

2.1 Linear model, least squares method and normal equations

■ Definition of a linear model

Y : dependent (response) variables

$X_1, \dots, X_p$: independent (explanatory) variables

$$\Rightarrow Y = b_1X_1 + b_2X_2 + \dots + b_pX_p + e ; e \sim (0, \sigma^2)$$

Note: In general, $X_1 = 1$

- Regression analysis (a case of n observations)

$$y_i = b_1X_{i1} + b_2X_{i2} + \dots + b_pX_{ip} + e_i ; i = 1, 2, \dots, n$$

$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} x_{11} & x_{12} & \dots & x_{1p} \\ \vdots & \vdots & & \vdots \\ x_{n1} & x_{n2} & \dots & x_{np} \end{pmatrix} \begin{pmatrix} b_1 \\ \vdots \\ b_p \end{pmatrix} + \begin{pmatrix} e_1 \\ \vdots \\ e_n \end{pmatrix}$$

$$\Leftrightarrow \underline{y} = X\underline{b} + \underline{e} \quad (X : \text{incidence matrix or design matrix}): \text{Matrix equation}$$

Example: Aerobic Fitness

	age	weight	runtime	rst pulse	run pulse	max pulse	oxy
1	44	89.47	11.37	62	178	182	44.61
2	40	75.07	10.07	62	185	185	45.31
3	44	85.84	8.65	45	156	168	54.30
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮
29	49	76.32	9.4	56	186	188	48.67
30	48	61.24	11.5	52	170	176	47.92
31	52	82.78	10.5	53	170	172	47.47

* The number of observations : $n = 31$

* The number of variables : $p = 7$

* Dependent variable(1) : oxy

* Independent variables(6) : age, weight, runtime, rstpulse, runpulse, maxpulse

* Regression analysis : $y_i = b_1 X_{i1} + b_2 X_{i2} + \dots + b_p X_{ip} + e_i$; $i = 1, 2, \dots, n$

$$\Rightarrow \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} x_{11} & x_{12} & \dots & x_{1p} \\ \vdots & \vdots & & \vdots \\ x_{n1} & x_{n2} & \dots & x_{np} \end{pmatrix} \begin{pmatrix} b_1 \\ \vdots \\ b_p \end{pmatrix} + \begin{pmatrix} e_1 \\ \vdots \\ e_n \end{pmatrix} \Leftrightarrow \underline{y} = X\underline{b} + \underline{e}$$

$$\Rightarrow \begin{pmatrix} 44.61 \\ \vdots \\ 47.47 \end{pmatrix} = \begin{pmatrix} 1 & 44 & \dots & 182 \\ \vdots & \vdots & & \vdots \\ 1 & 52 & \dots & 172 \end{pmatrix} \begin{pmatrix} b_1 \\ \vdots \\ b_p \end{pmatrix} + \begin{pmatrix} e_1 \\ \vdots \\ e_n \end{pmatrix}$$

$$\therefore \hat{\underline{b}} = (X^t X)^{-1} X^t \underline{y} \quad \text{if } \exists (X^t X)^{-1}$$

$$\Rightarrow y_i = 102.9153 X_{i1} - 0.2267 X_{i2} + \dots + 0.3037 X_{i7} + e_i ; i = 1, 2, \dots, 31$$

• Analysis of variance (ANOVA)

$$y_{ij} = \mu + \alpha_i + e_{ij}$$

Example: weights of 6 plants ($i = 1, 2, 3$ $j = 1, \dots, n_i$)

Normal	Off-type	Aberrant
$y_{11} = 101$	$y_{21} = 84$	$y_{31} = 32$
$y_{12} = 105$	$y_{22} = 88$	
$y_{13} = 94$		

$$\Rightarrow \begin{array}{lll} 101 = y_{11} = \mu + \alpha_1 & & + e_{11} \\ 105 = y_{12} = \mu + \alpha_1 & & + e_{12} \\ 94 = y_{13} = \mu & + \alpha_1 & + e_{13} \\ 84 = y_{21} = \mu & + \alpha_2 & + e_{21} \\ 88 = y_{22} = \mu & + \alpha_2 & + e_{22} \\ 32 = y_{31} = \mu & & + \alpha_3 + e_{31} \end{array} \Leftrightarrow \begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \\ y_{31} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{pmatrix} + \begin{pmatrix} e_{11} \\ e_{12} \\ e_{13} \\ e_{21} \\ e_{22} \\ e_{31} \end{pmatrix} \Leftrightarrow \underline{y} = X\underline{b} + \underline{e}$$

$$\therefore \hat{\underline{b}} = (X^t X)^{-1} X^t \underline{y} \quad \text{if } \exists (X^t X)^{-1} \quad \dots\dots\dots (1)$$

■ Least squares method(LSM)

For

$$y_i = b_1 X_{i1} + b_2 X_{i2} + \cdots + b_p X_{ip} + e_i \ ; \ i = 1, 2, \dots, n \ (\ n \geq p \) \Leftrightarrow \underline{y} = X\underline{b} + \underline{e}$$

$$S = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i - f(x_i, b))^2 = \underline{e}^t \underline{e} = (\underline{y} - X\underline{b})^t (\underline{y} - X\underline{b}) : \text{Sum of squared residuals}$$

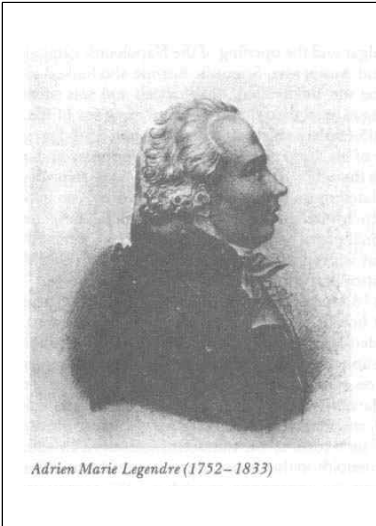
To find parameter values $b_1, b_2, \dots, b_p$ for which the model "best" fits data,
the least squares method defines "best" as when the sum S of squared residuals is a minimum

i.e.

$$\text{Min}_{b_1, \dots, b_p} \sum_{i=1}^n (y_i - b_1 x_{i1} - \cdots - b_p x_{ip})^2 \Leftrightarrow \text{Min}_{\underline{b}} (\underline{y} - X\underline{b})^t (\underline{y} - X\underline{b})$$

: The least squares method

History of Least Squares Method



Adrien-Marie Legendre (born 18 September 1752 – died 10 January 1833)

In 1805, the idea of least squares analysis was formulated by the Legendre.

- In 1805, he published the work by which he is chiefly known in the history of statistics:

Nouvelles methodes pour la determination des orbites des cometes.
– *Sur la methods des moindres quarres*

<혜성궤도를 결정하기 위한 새로운 방법> -부록 (최소제곱부분에 대하여)

- Ten years after Legendre's 1805 appendix, the method of LS was a standard tool in astronomy and geodesy in France, Italy, and Prussia. By 1825 the same was true in England.
- LSM was closely associated with three of the major scientific problems of 18c:
 - i) to determine and represent mathematically the motion of the moon
 - ii) Motions of the planets Jupiter and Saturn
 - iii) to determine the shape of figure of the earth.

Note : Matrix Derivatives(Differentiation)

$$\forall \underline{a}, \underline{v}, \underline{w} \in R^p$$

i) Scalar

$$- \frac{\partial}{\partial \underline{v}} \underline{a}^t \underline{v} = \frac{\partial}{\partial \underline{v}} \underline{v}^t \underline{a} = \begin{pmatrix} \frac{\partial}{\partial v_1} \underline{a}^t \underline{v} \\ \vdots \\ \frac{\partial}{\partial v_p} \underline{a}^t \underline{v} \end{pmatrix} = \begin{pmatrix} a_1 \\ \vdots \\ a_p \end{pmatrix} = \underline{a} : \text{round } \underline{a}^t \underline{v} \text{ over round } \underline{v} \text{ (derivative of } \underline{a}^t \underline{v} \text{ wrt } \underline{v})$$

$$- \frac{\partial \underline{a}^t \underline{v}}{\partial \underline{v}^t} = \left(\frac{\partial \underline{a}^t \underline{v}}{\partial v_1}, \dots, \frac{\partial \underline{a}^t \underline{v}}{\partial v_p} \right) = (a_1, \dots, a_p) = \underline{a}^t$$

ii) Vectors

$$A_{p \times p} = (\underline{a}_1, \underline{a}_2, \dots, \underline{a}_p) \quad \text{and} \quad \underline{v}^t A = (\underline{v}^t \underline{a}_1, \underline{v}^t \underline{a}_2, \dots, \underline{v}^t \underline{a}_p)$$

$$- \frac{\partial \underline{v}^t A}{\partial \underline{v}} = \left(\frac{\partial \underline{v}^t \underline{a}_1}{\partial \underline{v}}, \frac{\partial \underline{v}^t \underline{a}_2}{\partial \underline{v}}, \dots, \frac{\partial \underline{v}^t \underline{a}_p}{\partial \underline{v}} \right) = (\underline{a}_1, \underline{a}_2, \dots, \underline{a}_p) = A$$

$$- \left(\frac{\partial}{\partial \underline{v}^t} (A \underline{v}) \right)^t = \frac{\partial}{\partial \underline{v}} (A \underline{v})^t = \frac{\partial}{\partial \underline{v}} (\underline{v}^t A^t) = A^t$$

iii) Quadratic forms

$$\forall \underline{v}, \underline{w} \in R^p, \quad A_{p \times p} = (\underline{a}_1, \underline{a}_2, \dots, \underline{a}_p)$$

$$- \frac{\partial \underline{v}^t A \underline{v}}{\partial \underline{v}} = (A + A^t) \underline{v}$$

$$- \frac{\partial \underline{v}^t A \underline{v}}{\partial \underline{v}} = 2A \underline{v} \quad \text{if } A = A^t$$

$$- \frac{\partial \underline{v}^t A \underline{w}}{\partial \underline{v}} = A \underline{w}$$

$$- \frac{\partial \underline{v}^t \underline{v}}{\partial \underline{v}} = 2 \underline{v}$$

Proof of equation (1): $\hat{\underline{b}} = (X^t X)^{-1} X^t \underline{y}$

- Regression analysis and analysis of variance (ANOVA)

$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} x_{11} & x_{12} & \cdots & x_{1p} \\ \vdots & \vdots & & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{np} \end{pmatrix} \begin{pmatrix} b_1 \\ \vdots \\ b_p \end{pmatrix} + \begin{pmatrix} e_1 \\ \vdots \\ e_n \end{pmatrix} \Rightarrow \underline{y} = X\underline{b} + \underline{e} \Leftarrow \begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \\ y_{31} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{pmatrix} + \begin{pmatrix} e_{11} \\ e_{12} \\ e_{13} \\ e_{21} \\ e_{22} \\ e_{31} \end{pmatrix}$$

i) Matrix proof

$$S = \underline{e}^t \underline{e} = (\underline{y} - X\underline{b})^t (\underline{y} - X\underline{b})$$

$$= \underline{y}^t \underline{y} - \underline{b}^t X^t \underline{y} - \underline{y}^t X \underline{b} + \underline{b}^t X^t X \underline{b}$$

$$= \underline{y}^t \underline{y} - 2\underline{b}^t X^t \underline{y} + \underline{b}^t X^t X \underline{b}$$

$$\frac{\partial S}{\partial \underline{b}} = -2X^t \underline{y} + 2X^t X \underline{b} = 0 \quad \Leftrightarrow \quad X^t X \underline{b} = X^t \underline{y} : \text{normal equation}$$

$$\Leftrightarrow \hat{\underline{b}} = (X^t X)^{-1} X^t \underline{y}$$

($\because$ rank(X)=full column rank = $p \Rightarrow \exists (X^t X)^{-1}$)
: least square solution for $\underline{b}$ = the unique least squares estimator

ii) Elementwise proof

$$y_i = b_1 x_{i1} + b_2 x_{i2} + \dots + b_p x_{ip} + e_i, \quad i = 1, 2, \dots, n$$

$$S = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i - b_1 x_{i1} - \dots - b_p x_{ip})^2 : \text{sum of squared errors}$$

$$\Rightarrow \min_{b_1, \dots, b_p} S$$

$$f(b_1, b_2, \dots, b_p) = \sum_{i=1}^n (y_i - b_1 x_{i1} - \dots - b_p x_{ip})^2 : \text{objective function}$$

$$\left\{ \begin{array}{l} \frac{\partial f}{\partial b_1} = \sum_{i=1}^n 2(y_i - b_1 x_{i1} - \dots - b_p x_{ip})(-x_{i1}) = 0 \\ \frac{\partial f}{\partial b_2} = \sum_{i=1}^n 2(y_i - b_1 x_{i1} - \dots - b_p x_{ip})(-x_{i2}) = 0 \\ \vdots \\ \frac{\partial f}{\partial b_p} = \sum_{i=1}^n 2(y_i - b_1 x_{i1} - \dots - b_p x_{ip})(-x_{ip}) = 0 \end{array} \right. : \text{partial derivative of } f(\bullet) \text{ w.r.t } b_j$$

$$\Leftrightarrow \begin{cases} \sum_{i=1}^n (b_1 x_{i1} x_{i1} + \dots + b_p x_{i1} x_{ip}) = \sum_{i=1}^n x_{i1} y_i \\ \sum_{i=1}^n (b_1 x_{i2} x_{i1} + \dots + b_p x_{i2} x_{ip}) = \sum_{i=1}^n x_{i2} y_i \\ \vdots \\ \sum_{i=1}^n (b_1 x_{ip} x_{i1} + \dots + b_p x_{ip} x_{ip}) = \sum_{i=1}^n x_{ip} y_i \end{cases} \Leftrightarrow \begin{cases} b_1 \sum_{i=1}^n x_{i1} x_{i1} + \dots + b_p \sum_{i=1}^n x_{i1} x_{ip} = \sum_{i=1}^n x_{i1} y_i \\ b_1 \sum_{i=1}^n x_{i2} x_{i1} + \dots + b_p \sum_{i=1}^n x_{i2} x_{ip} = \sum_{i=1}^n x_{i2} y_i \\ \vdots \\ b_1 \sum_{i=1}^n x_{ip} x_{i1} + \dots + b_p \sum_{i=1}^n x_{ip} x_{ip} = \sum_{i=1}^n x_{ip} y_i \end{cases}$$

$$\Leftrightarrow \begin{cases} b_1 a_{11} + \dots + b_p a_{1p} = c_1 \\ b_1 a_{21} + \dots + b_p a_{2p} = c_2 \\ \vdots \\ b_1 a_{p1} + \dots + b_p a_{pp} = c_p \end{cases} \quad (\because \sum_{i=1}^n x_{ij} x_{ik} = a_{jk}, j, k = 1, \dots, p)$$

$$\Leftrightarrow \begin{pmatrix} a_{11} & \dots & a_{1p} \\ \vdots & & \vdots \\ a_{p1} & \dots & a_{pp} \end{pmatrix} \begin{pmatrix} b_1 \\ \vdots \\ b_p \end{pmatrix} = \begin{pmatrix} c_1 \\ \vdots \\ c_p \end{pmatrix}$$

$$\Leftrightarrow A \underline{b} = \underline{c} : \text{normal equation}$$

$$\text{In fact, } \begin{cases} A = X^t X \\ \underline{c} = X^t \underline{y} \end{cases} \Rightarrow X^t X \underline{b} = X^t \underline{y} : \text{normal equation}$$

2.2 Solution of (simultaneous) equations and inverse matrices

$$\begin{cases} 2b_1 - 5b_2 + 4b_3 = -3 \\ b_1 - 2b_2 + b_3 = 5 \\ b_1 - 4b_2 + 6b_3 = 10 \end{cases} \Rightarrow \begin{cases} b_1 = 124 \\ b_2 = 75 \\ b_3 = 31 \end{cases} \quad (\text{by Gaussian elimination method})$$

$$\Leftrightarrow \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 6 \end{pmatrix} \begin{pmatrix} b_1 \\ b_2 \\ b_3 \end{pmatrix} = \begin{pmatrix} -3 \\ 5 \\ 10 \end{pmatrix} : \text{Matrix equation}$$

$$\Leftrightarrow A\underline{b} = \underline{c}$$

$$\exists A^{-1} \Rightarrow A^{-1}A\underline{b} = A^{-1}\underline{c} \Rightarrow \hat{\underline{b}} = A^{-1}\underline{c} = \begin{pmatrix} 124 \\ 75 \\ 31 \end{pmatrix}$$

■ Definition of identity matrix

$$I = \text{diag}(1, 1, \dots, 1)$$

: An identity matrix

having all diagonal elements equal to 1

For any matrices A and B

$$AI = A, IB = B$$

■ Definition of inverse matrix

For square matrix A , $\exists$ square matrix B s.t.

$$AB = BA = I$$

$$\Rightarrow B = A^{-1} : \text{The inverse of } A$$

Remark: simple properties of inverse matrix

- $(AB)^{-1} = B^{-1}A^{-1}$
- $(A^t)^{-1} = (A^{-1})^t$
- $(A^{-1})^{-1} = A$

since $AB = I$,

$$AB = I \rightarrow A^{-1}AB = A^{-1}I \rightarrow IB = A^{-1}$$

$$\text{Note : } A \rightarrow I, I \rightarrow A^{-1}$$

$$\Rightarrow (A|I) \rightarrow (I|A^{-1})$$

$$\left(\begin{array}{ccc|ccc} 2 & -5 & 4 & 1 & 0 & 0 \\ 1 & -2 & 1 & 0 & 1 & 0 \\ 1 & -4 & 6 & 0 & 0 & 1 \end{array} \right) \rightarrow \left(\begin{array}{ccc|ccc} 1 & 0 & 0 & -8 & 14 & 3 \\ 0 & 1 & 0 & -5 & 8 & 2 \\ 0 & 0 & 1 & -2 & 3 & 1 \end{array} \right) : \text{Augmented matrix}$$

- Gauss-Jordan elimination method

$$\text{M1) } A\underline{b} = \underline{c}$$

$$\exists A^{-1} \Rightarrow A^{-1}A\underline{b} = A^{-1}\underline{c} \Leftrightarrow I\underline{b} = A^{-1}\underline{c}$$

$$\text{Note : } A \rightarrow I, \underline{c} \rightarrow A^{-1}\underline{c}$$

$$\Rightarrow (A|\underline{c}) \rightarrow (I|A^{-1}\underline{c})$$

$$\left(\begin{array}{ccc|c} 2 & -5 & 4 & -3 \\ 1 & -2 & 1 & 5 \\ 1 & -4 & 6 & 10 \end{array} \right) \rightarrow \left(\begin{array}{ccc|c} 1 & 0 & 0 & 124 \\ 0 & 1 & 0 & 75 \\ 0 & 0 & 1 & 31 \end{array} \right)$$

M2) $A\underline{b} = \underline{c}$

$$\exists A^{-1} \Rightarrow A\underline{b} = I\underline{c} \Leftrightarrow I\underline{b} = A^{-1}\underline{c}$$

Note : $A \rightarrow I$, $I \rightarrow A^{-1}$, $\underline{c} \rightarrow A^{-1}\underline{c}$

$$\Rightarrow (A|\underline{c}|I) \rightarrow (I|A^{-1}\underline{c}|A^{-1})$$

$$\left(\begin{array}{ccc|c|ccc} 2 & -5 & 4 & -3 & 1 & 0 & 0 \\ 1 & -2 & 1 & 5 & 0 & 1 & 0 \\ 1 & -4 & 6 & 10 & 0 & 0 & 1 \end{array} \right) \rightarrow \left(\begin{array}{ccc|c|ccc} 1 & 0 & 0 & 124 & -8 & 14 & 3 \\ 0 & 1 & 0 & 75 & -5 & 8 & 2 \\ 0 & 0 & 1 & 31 & -2 & 3 & 1 \end{array} \right)$$

Summary : Elimination Methods

- $(A|I) \rightarrow (I|A^{-1})$
- $(A|\underline{c}) \rightarrow (I|A^{-1}\underline{c})$
- $(A|\underline{c}|I) \rightarrow (I|A^{-1}\underline{c}|A^{-1})$

Home work : Do it !

2.3 Conditions for existence of the inverse

■ Conditions for the existence of A^{-1}

- i) $A_{n \times n}$: non-singular
- ii) $|A| \neq 0$
- iii) $rank(A) = n$: full-rank
- iv) A has n linearly independent rows
- v) A has n linearly independent columns
- vi) $A\underline{x} = \underline{0}$ has the sole solution $\underline{x} = \underline{0}$
- vii) Number of eigenvalues of $A = n$

■ Properties of A^{-1}

- $A^{-1}A = AA^{-1} = I$ (commutative law)
- $\exists ! A^{-1}$ (uniqueness)
- $|A^{-1}| = \frac{1}{|A|}$
- $|A^{-1}| \neq 0$
- $(A^{-1})^{-1} = A$
- $(A^t)^{-1} = (A^{-1})^t$
- $A^t = A \Rightarrow (A^{-1})^t = A^{-1}$
- $(AB)^{-1} = B^{-1}A^{-1}$
- $(A+B)^{-1} = A^{-1} - A^{-1}(B^{-1} + A^{-1})A^{-1}$ ($\Leftarrow A, B$ and $A+B$: nonsingular)

2.4 Calculation of the inverse in R and SAS/IML

$A\underline{b} = \underline{c} \Rightarrow \underline{b} = A^{-1}\underline{c}$	
R	SAS/IML
• solve(A) <code>b<-solve(A)%*%c (b<- solve(A, c))</code>	• inv(A) <code>b=inv(A)*c;</code>

Refer to the text p. 40

2.5 Linear dependence and independence

$$A_1 = \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 5 \end{pmatrix} \Rightarrow |A_1| = \begin{vmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 5 \end{vmatrix}$$

$$= [2(-2)5 + 1(-4)4 + 1(-5)1] - [4(-2)1 + 1(-5)5 + 2(-4)1]$$

$$= [-20 - 16 - 5] - [-8 - 25 - 8] = -25 - 16 + 16 + 25 = 0 \quad (\because \text{Rule of Sarrus})$$

$$\Rightarrow \nexists A^{-1} \Leftrightarrow A : \text{singular}$$

Note

Rule of Sarrus or Sarrus' Rule is a method and a memorization scheme to compute the determinant of a 3×3 matrix. It is named after the French mathematician Pierre Frédéric Sarrus.



Sarrus(1798-1861)

- Rows

$$2 \begin{pmatrix} 2 \\ -5 \\ 4 \end{pmatrix} - 3 \begin{pmatrix} 1 \\ -2 \\ 1 \end{pmatrix} = \begin{pmatrix} 1 \\ -4 \\ 5 \end{pmatrix} : \text{linear dependence}$$

- Columns

$$-3 \begin{pmatrix} 2 \\ 1 \\ 1 \end{pmatrix} - 2 \begin{pmatrix} -5 \\ -2 \\ -4 \end{pmatrix} = \begin{pmatrix} 4 \\ 1 \\ 5 \end{pmatrix} : \text{linear dependence}$$

■ LIN (linear independence or linearly independent)

$$X_{n \times n} \underline{k} = (\underline{x}_1, \underline{x}_2, \dots, \underline{x}_n) \begin{pmatrix} k_1 \\ k_2 \\ \vdots \\ k_n \end{pmatrix} = \sum_{i=1}^n k_i \underline{x}_i \quad : \text{linear combination}$$

$$X \underline{k} = \underline{0} \Rightarrow \underline{k} = \underline{0}$$

$$\Leftrightarrow X \underline{k} = \underline{0} \text{ only for } \underline{k} = \underline{0}$$

: The columns of X are LIN vectors

■ LDN (linear dependence or linearly dependent)

$$X_{n \times n} \underline{k} = (\underline{x}_1, \underline{x}_2, \dots, \underline{x}_n) \begin{pmatrix} k_1 \\ k_2 \\ \vdots \\ k_n \end{pmatrix} = \sum_{i=1}^n k_i \underline{x}_i \quad : \text{linear combination}$$

$$X \underline{k} = \underline{0} \Rightarrow \underline{k} \neq \underline{0}$$

$$\Leftrightarrow X \underline{k} = \underline{0} \text{ for some non-null } \underline{k} \text{ } (\underline{k} \neq \underline{0})$$

: The columns of X are LDN vectors

Example

$$\text{i) } 2\begin{pmatrix} 2 \\ -5 \\ 4 \end{pmatrix} - 3\begin{pmatrix} 1 \\ -2 \\ 1 \end{pmatrix} - 1\begin{pmatrix} 1 \\ -4 \\ 5 \end{pmatrix} = \underline{0} \Leftrightarrow \begin{pmatrix} 2 & 1 & 1 \\ -5 & -2 & -4 \\ 4 & 1 & 5 \end{pmatrix} \begin{pmatrix} 2 \\ -3 \\ 1 \end{pmatrix} = \underline{0}$$

$$\Leftrightarrow A_1 \underline{k} = \underline{0}$$

$$\Rightarrow \underline{k} = \begin{pmatrix} 2 \\ -3 \\ -1 \end{pmatrix} \neq \underline{0} : \text{LDN,}$$

Note: $|A_1| = 0$

$$\text{ii) } A_2 = \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 6 \end{pmatrix}$$

$$k_1 \begin{pmatrix} 2 \\ 1 \\ 1 \end{pmatrix} + k_2 \begin{pmatrix} -5 \\ -2 \\ -4 \end{pmatrix} + k_3 \begin{pmatrix} 4 \\ 1 \\ 6 \end{pmatrix} = \underline{0} \Leftrightarrow \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 6 \end{pmatrix} \begin{pmatrix} k_1 \\ k_2 \\ k_3 \end{pmatrix} = \underline{0} \Leftrightarrow A_2 \underline{k} = \underline{0}$$

Note: $|A_2| = 17 \neq 0$

$$\exists A_2^{-1} \Rightarrow A_2^{-1} A_2 \underline{k} = A_2^{-1} \underline{0}$$

$$\Leftrightarrow \underline{k} = \underline{0} \therefore \text{The columns of } A_2 \text{ are LIN}$$

■ Definition of rank

Rank of $A \equiv rank(A) \equiv r(A)$

: The number of linearly independent rows (and columns) in A

Example

$$A_1 = \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 5 \end{pmatrix}, \quad A_2 = \begin{pmatrix} 2 & -5 & 4 \\ 1 & -2 & 1 \\ 1 & -4 & 6 \end{pmatrix}$$

$$rank(A_1) = 2 \Leftrightarrow \det(A_1) = 0$$

$$rank(A_2) = 3 \Leftrightarrow \det(A_2) \neq 0$$

■ Properties of $rank(A_{m \times n})$

- i) $rank(A_{m \times n}) = \min(m, n) \Rightarrow A$: full rank
- ii) $rank(A_{m \times n}) = m \Rightarrow A$: full row rank
- iii) $rank(A_{m \times n}) = n \Rightarrow A$: full column rank

2.6 Determinants

■ Application to statistics

i) Variance

$$X_{n \times p} = (x_{ij}) \quad , \quad i = 1, \dots, n \quad , \quad j = 1, \dots, p$$

: Data matrix

$$\rightarrow Y = (x_{ij} - \bar{x}_j)$$

: Centering matrix

$$Y^t Y = \begin{pmatrix} \sum_{i=1}^n (x_{i1} - \bar{x}_1)^2 & \sum_{i=1}^n (x_{i1} - \bar{x}_1)(x_{i2} - \bar{x}_2) & \cdots & \sum_{i=1}^n (x_{i1} - \bar{x}_1)(x_{ip} - \bar{x}_p) \\ & \sum_{i=1}^n (x_{i2} - \bar{x}_2)^2 & \cdots & \sum_{i=1}^n (x_{i2} - \bar{x}_2)(x_{ip} - \bar{x}_p) \\ & & \ddots & \vdots \\ sym & & & \sum_{i=1}^n (x_{ip} - \bar{x}_p)^2 \end{pmatrix}$$

$$= (n-1)S = (n-1) \begin{pmatrix} s_{11} & s_{12} & \cdots & s_{1p} \\ & s_{22} & \cdots & s_{2p} \\ & & \ddots & \vdots \\ sym & & & s_{pp} \end{pmatrix}$$

(S : The variance-covariance matrix)

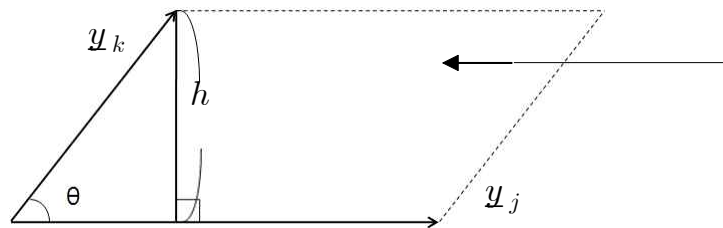
- Variance

- Generalized : $|S|$
- Total : $tr(S) = s_{11} + \cdots + s_{pp}$

- Geometric view of $|S|$, with $S = \begin{pmatrix} S_{kk} & S_{kj} \\ S_{jk} & S_{jj} \end{pmatrix}$ and $p = 2$

$$\underline{y}_k = (x_{1k} - \bar{x}_k, \cdots, x_{nk} - \bar{x}_k)^t$$

$$\underline{y}_j = (x_{1j} - \bar{x}_j, \cdots, x_{nj} - \bar{x}_j)^t$$



$$Area = \| \underline{y}_k \| \times \| \underline{y}_j \| \times \sin \theta$$

$$= \sqrt{\sum_{i=1}^n (x_{ik} - \bar{x}_k)^2} \sqrt{\sum_{i=1}^n (x_{ij} - \bar{x}_j)^2} \times \sqrt{1 - \cos^2 \theta}$$

Note

$$\bullet \sum_{i=1}^n (x_{ik} - \bar{x}_k)^2 = (n-1)s_{kk}$$

$$\bullet \sum_{i=1}^n (x_{ij} - \bar{x}_j)^2 = (n-1)s_{jj}$$

$$\begin{aligned} \bullet \cos\theta &= \frac{\underline{y}_k^t \underline{y}_j}{\|\underline{y}_k\| \|\underline{y}_j\|} = \frac{\sum_{i=1}^n (x_{ik} - \bar{x}_k)(x_{ij} - \bar{x}_j)}{\sqrt{\sum_{i=1}^n (x_{ik} - \bar{x}_k)^2} \sqrt{\sum_{i=1}^n (x_{ij} - \bar{x}_j)^2}} \\ &= \frac{(n-1)s_{kj}}{\sqrt{(n-1)s_{kk}} \sqrt{(n-1)s_{jj}}} \\ &= \frac{s_{kj}}{\sqrt{s_{kk}} \sqrt{s_{jj}}} = r_{kj} \end{aligned}$$

$$\therefore \text{Area} = (n-1) \sqrt{s_{kk}} \sqrt{s_{jj}} \sqrt{1 - r_{kj}^2}$$

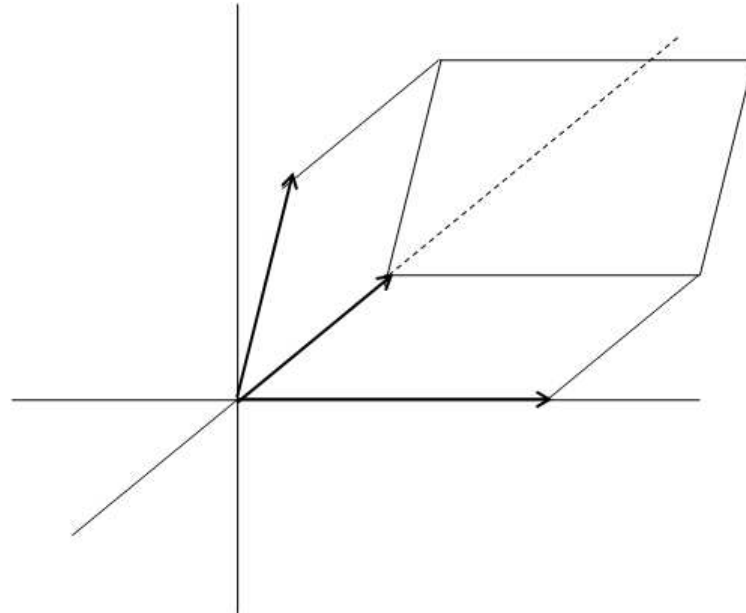
$$|S| = \begin{vmatrix} s_{kk} & s_{kj} \\ s_{jk} & s_{jj} \end{vmatrix} = s_{kk} s_{jj} - s_{kj} s_{jk}$$

$$= s_{kk} s_{jj} - (r_{jk} \sqrt{s_{kk}} \sqrt{s_{jj}})(r_{jk} \sqrt{s_{kk}} \sqrt{s_{jj}}) \quad (\because r_{kj} = \frac{s_{kj}}{\sqrt{s_{kk}} \sqrt{s_{jj}}} = r_{jk})$$

$$= s_{kk} s_{jj} (1 - r_{jk}^2)$$

$$\therefore \text{ If } p = 2, \quad |S| = \frac{(\text{Area})^2}{(n-1)^2}$$

$$- \text{ If } p > 2, \quad |S| = \frac{(\text{Volume})^2}{(n-1)^p}$$



ii) Multivariate normal distribution

$$x \sim N(\mu, \sigma^2), \sigma > 0$$

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$$= (2\pi)^{-1/2} (\sigma^2)^{-1/2} \exp\left(-\frac{1}{2}(x-\mu)(\sigma^2)^{-1}(x-\mu)\right)$$

$$\Rightarrow \underline{x} = (x_1, \dots, x_p)^t \sim N_p(\underline{\mu}, \Sigma) \quad , \quad \Sigma > 0$$

$$f(\underline{x}) = (2\pi)^{-p/2} |\Sigma|^{-1/2} \exp\left(-\frac{1}{2}(\underline{x}-\underline{\mu})^t \Sigma^{-1}(\underline{x}-\underline{\mu})\right)$$

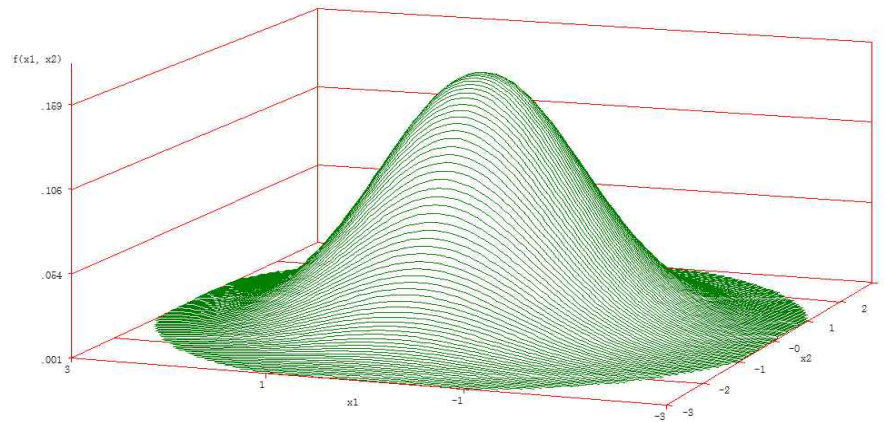
↳ Ellipsoid

- $p = 2$, Bivariate normal

$$\underline{x} = (x_1, x_2) \sim N_2(\underline{\mu}, \Sigma) \quad \text{where } \underline{\mu} = (\mu_1, \mu_2)^t, \quad \Sigma = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix} = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$

$$|\Sigma| = \sigma_1^2\sigma_2^2 - \rho^2\sigma_1^2\sigma_2^2 = (1 - \rho^2)\sigma_1^2\sigma_2^2$$

$$\Sigma^{-1} = \frac{1}{|\Sigma|} \begin{pmatrix} \sigma_2^2 & -\rho\sigma_1\sigma_2 \\ -\rho\sigma_1\sigma_2 & \sigma_1^2 \end{pmatrix}$$



$$(x_1 - \mu_1 \quad x_2 - \mu_2) \Sigma^{-1} \begin{pmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{pmatrix} = \frac{1}{1 - \rho^2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]$$

$$\therefore f(\underline{x}) = f(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp(-q/2)$$

$$\text{where } q = \frac{1}{1 - \rho^2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]$$

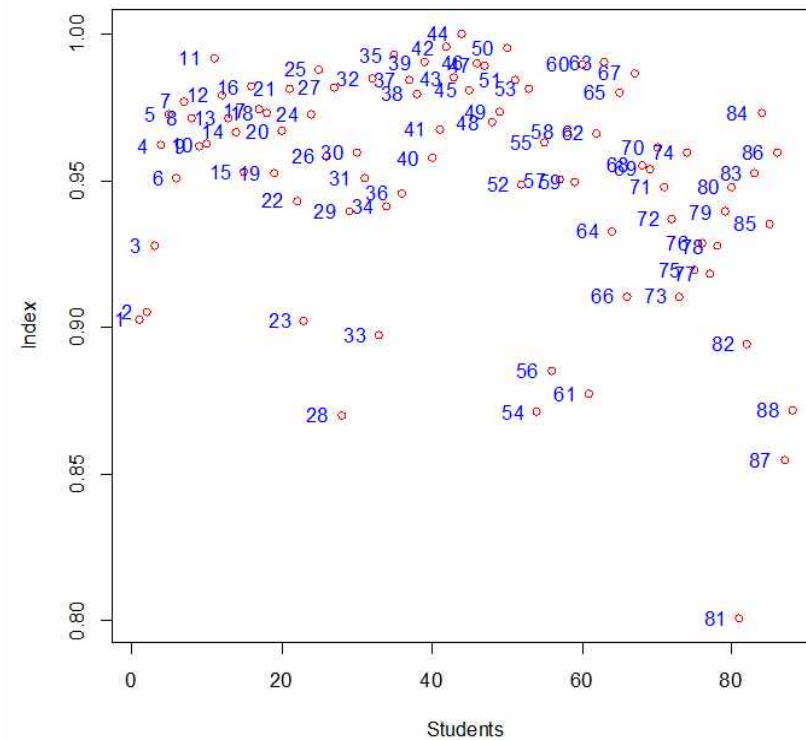
$$- \rho = 0 \quad \rightarrow \quad f(x_1, x_2) = f(x_1)f(x_2)$$

iii) Outlier

- $r_i^2 = \frac{|S_{(i)}|}{|S|}$: for single outlier where $S_{(i)}$ covariance matrix with x_i omitted.
- $\frac{|S_{(\underline{i})}|}{|S|}$: for multiple outlier where $\underline{i} = (i_1, i_2, \dots, i_t)$, $t = (\text{number of observations})$

Criterion: A relatively small value of r_i^2 indice x_i is potential outlier.

single outlier					
[1,]	0.9472527	[26,]	1.0055308	[56,]	0.9289727
		[27,]	1.0300230	[57,]	0.9970349
[2,]	0.9500749	[28,]	0.9129172	[58,]	1.0140230
[3,]	0.9736236	[29,]	0.9858263	[59,]	0.9962962
[4,]	1.0093201	[30,]	1.0069715	[60,]	1.0380680
[5,]	1.0204481	[31,]	0.9978288	[61,]	0.9208771
[6,]	0.9976272	[32,]	1.0333239	[62,]	1.0137424
[7,]	1.0250582	[33,]	0.9416272	[63,]	1.0389501
[8,]	1.0192968	[34,]	0.9877500	[64,]	0.9787027
[9,]	1.0092410	[35,]	1.0416638	[65,]	1.0281443
		[36,]	0.9924048	[66,]	0.9554624
[10,]	1.0097776	[37,]	1.0326007	[67,]	1.0350832
[11,]	1.0404495	[38,]	1.0278448	[68,]	1.0021611
[12,]	1.0273055	[39,]	1.0389238	[69,]	1.0010168
[13,]	1.0189797	[40,]	1.0047766	[70,]	1.0087386
[14,]	1.0142306	[41,]	1.0150475	[71,]	0.9946434
[15,]	0.9998236	[42,]	1.0444162	[72,]	0.9832537
[16,]	1.0302959	[43,]	1.0335075	[73,]	0.9554405
[17,]	1.0221134	[44,]	1.0493268	[74,]	1.0068274
[18,]	1.0208078	[45,]	1.0290247	[75,]	0.9647899
[19,]	0.9996248	[46,]	1.0385267	[76,]	0.9746207
[20,]	1.0143645	[47,]	1.0379448	[77,]	0.9633846
[21,]	1.0294478	[48,]	1.0175129	[78,]	0.9733431
[22,]	0.9894279	[49,]	1.0214440	[79,]	0.9856481
[23,]	0.9466385	[50,]	1.0440324	[80,]	0.9946146
[24,]	1.0203104	[51,]	1.0327465	[81,]	0.8404648
[25,]	1.0365367	[52,]	0.9954448	[82,]	0.9384063
		[53,]	1.0297299	[83,]	0.9996551
		[54,]	0.9141544	[84,]	1.0210044
		[55,]	1.0104093	[85,]	0.9811699
				[86,]	1.0066482
				[87,]	0.8970167
				[88,]	0.9146625



■ Axiomatic definition of determinant :

$$A = \begin{pmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{pmatrix} = \begin{bmatrix} \underline{a_1^t} \\ \underline{a_2^t} \end{bmatrix}$$

Axiom i) Factorization

$$\begin{aligned} \begin{vmatrix} ka_{11} & ka_{12} \\ a_{21} & a_{22} \end{vmatrix} &= ka_{11}a_{22} - ka_{12}a_{21} = k(a_{11}a_{22} - a_{12}a_{21}) = k \begin{vmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{vmatrix} \\ \Leftrightarrow \begin{vmatrix} k\underline{a_1^t} \\ \underline{a_2^t} \end{vmatrix} &= \det \begin{bmatrix} k\underline{a_1^t} \\ \underline{a_2^t} \end{bmatrix} = k \det \begin{bmatrix} \underline{a_1^t} \\ \underline{a_2^t} \end{bmatrix} \Rightarrow \det \begin{bmatrix} \underline{a_1^t} \\ \vdots \\ k\underline{a_i^t} \\ \vdots \\ \underline{a_p^t} \end{bmatrix} = k \det \begin{bmatrix} \underline{a_1^t} \\ \vdots \\ \underline{a_i^t} \\ \vdots \\ \underline{a_p^t} \end{bmatrix} \end{aligned}$$

Axiom ii) Adding multiples of a row(column) to a row(column)

$$\begin{aligned} \begin{vmatrix} a_{11} + k & a_{12} + k \\ a_{21} & a_{22} \end{vmatrix} &= (a_{11} + k)a_{22} - a_{21}(a_{12} + k) = a_{11}a_{22} - a_{21}a_{12} + ka_{22} - ka_{21} = \begin{vmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{vmatrix} + \begin{vmatrix} k & k \\ a_{21} & a_{22} \end{vmatrix} \\ \Leftrightarrow \det \begin{bmatrix} \underline{a_1^t} + k\underline{1_2^t} \\ \underline{a_2^t} \end{bmatrix} &= \det \begin{bmatrix} \underline{a_1^t} \\ \underline{a_2^t} \end{bmatrix} + \det \begin{bmatrix} k\underline{1_2^t} \\ \underline{a_2^t} \end{bmatrix} \Rightarrow \det \begin{bmatrix} \underline{a_1^t} \\ \vdots \\ \underline{a_i^t} + k\underline{1_p^t} \\ \vdots \\ \underline{a_p^t} \end{bmatrix} = \det \begin{bmatrix} \underline{a_1^t} \\ \vdots \\ \underline{a_i^t} \\ \vdots \\ \underline{a_p^t} \end{bmatrix} + \det \begin{bmatrix} \underline{a_1^t} \\ \vdots \\ k\underline{1_p^t} \\ \vdots \\ \underline{a_p^t} \end{bmatrix} \end{aligned}$$

Axiom iii) If two rows(columns) of A are the same, $|A| = 0$

$$\begin{vmatrix} a_{11} & a_{12} \\ a_{11} & a_{12} \end{vmatrix} = 0$$

Axiom iv) The determinant of the Identity matrix is 1

$$\begin{vmatrix} 1 & 0 \\ 0 & 1 \end{vmatrix} = 1$$

$$\Rightarrow \det(I_p) = \det(\underline{e}_1, \underline{e}_1, \dots, \underline{e}_p) = 1$$

Axiom v) For $kA_{p \times p}$, $\det(kA) = k^p \det(A)$

■ Zero determinants

$\det(A) \neq 0 \Leftrightarrow$ The vectors are linearly independent (LIN)

$$\Leftrightarrow A \text{ is non-singular} \Leftrightarrow \exists A^{-1}$$

2.7 Determinants and the theory of the inverse

■ Calculate the determinants

$$\bullet \quad \begin{vmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{vmatrix} = a_{11}a_{22} - a_{12}a_{21}$$

$$\bullet \quad |A| = \begin{vmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \\ a_{31} & a_{32} & a_{33} \end{vmatrix}$$

$$= a_{11}a_{22}a_{33} + a_{12}a_{23}a_{31} + a_{13}a_{21}a_{32} - (a_{13}a_{22}a_{31} + a_{12}a_{21}a_{33} + a_{11}a_{23}a_{32}) \quad : \text{Sarrus' Rule}$$

$$= a_{11} \begin{vmatrix} a_{22} & a_{23} \\ a_{32} & a_{33} \end{vmatrix} + a_{12}(-1) \begin{vmatrix} a_{21} & a_{23} \\ a_{31} & a_{33} \end{vmatrix} + a_{13} \begin{vmatrix} a_{21} & a_{22} \\ a_{31} & a_{32} \end{vmatrix} \quad : \text{Expansion by minors}$$

$$= a_{11}(-1)^{1+1}|M_{11}| + a_{12}(-1)^{1+2}|M_{12}| + a_{13}(-1)^{1+3}|M_{13}|$$

$$\begin{aligned}
 \bullet \quad |A_{n \times n}| &= \begin{vmatrix} a_{11} & a_{12} & \cdots & a_{1n} \\ \vdots & \vdots & & \vdots \\ a_{i1} & a_{i2} & \cdots & a_{in} \\ \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & \cdots & a_{nn} \end{vmatrix} = a_{i1}(-1)^{i+1}|M_{i1}| + a_{i2}(-1)^{i+2}|M_{i2}| + \cdots + a_{in}(-1)^{i+n}|M_{in}| \\
 &= \sum_{j=1}^n a_{ij}(-1)^{i+j}|M_{ij}|
 \end{aligned}$$

$$|A_{n \times n}| = \sum_{j=1}^n a_{ij}c_{ij}, \text{ for any } i$$

$$= \sum_{i=1}^n a_{ij}c_{ij}, \text{ for any } j$$

where $c_{ij} = (-1)^{i+j}|M_{ij}|$: Cofactor

: Cofactor Expansion

$$\bullet \quad 0 = \sum_{j=1}^n a_{ij}c_{kj}, \text{ for } i \neq k$$

$$\therefore \sum_{j=1}^3 a_{1j}c_{2j} = a_{11}(-1)^{2+1}|M_{21}| + a_{12}(-1)^{2+2}|M_{22}| + a_{13}(-1)^{2+3}|M_{23}|$$

$$= a_{11}(-1)^{2+1} \begin{vmatrix} a_{12} & a_{13} \\ a_{32} & a_{33} \end{vmatrix} + a_{12}(-1)^{2+2} \begin{vmatrix} a_{11} & a_{13} \\ a_{31} & a_{33} \end{vmatrix} + a_{13}(-1)^{2+3} \begin{vmatrix} a_{11} & a_{12} \\ a_{31} & a_{32} \end{vmatrix} = \begin{vmatrix} a_{11} & a_{12} & a_{13} \\ a_{11} & a_{12} & a_{13} \\ a_{31} & a_{32} & a_{33} \end{vmatrix} = 0$$

■ Derivation of the inverse by cofactors

Note

$$\left. \begin{aligned} A &= \begin{pmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{pmatrix} \\ |A| &= a_{11}a_{22} - a_{12}a_{21} \neq 0 \\ A^{-1} &= \frac{1}{|A|} \begin{pmatrix} a_{22} & -a_{12} \\ -a_{21} & a_{11} \end{pmatrix} \end{aligned} \right\} \Rightarrow A^{-1}A = I$$

$$\begin{aligned} \Leftrightarrow & \frac{1}{|A|} \begin{pmatrix} a_{22} & -a_{12} \\ -a_{21} & a_{11} \end{pmatrix} \begin{pmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{pmatrix} \\ &= \frac{1}{|A|} \begin{pmatrix} a_{11}a_{22} - a_{12}a_{21} & a_{22}a_{12} - a_{12}a_{22} \\ -a_{11}a_{21} + a_{11}a_{21} & -a_{21}a_{12} + a_{11}a_{22} \end{pmatrix} \\ &= \frac{1}{|A|} \begin{pmatrix} |A| & 0 \\ 0 & |A| \end{pmatrix} = \frac{|A|}{|A|} \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} = I \end{aligned}$$

- Adjugate of A

$$c_{ij} = (-1)^{i+j} |M_{ij}|$$

$$\Rightarrow \{c_{ij}\} = C$$

$$\Rightarrow \text{adj}A = C^t : \text{Adjugate (or adjoint) matrix of } A$$

Theorem

- $A_{n \times n}$ with $|A| \neq 0$

$$\Rightarrow A^{-1} = \frac{1}{|A|} C^t = \frac{1}{|A|} \text{adj}A$$

Theorem $AC^t = |A| \times I$

$$A^{-1} = \frac{1}{|A|} C^t$$

Proof) $C^t = \begin{pmatrix} c_{11} & \cdots & c_{1n} \\ \vdots & & \vdots \\ c_{n1} & \cdots & c_{nn} \end{pmatrix}^t = \begin{pmatrix} c_{11} & \cdots & c_{n1} \\ \vdots & & \vdots \\ c_{1n} & \cdots & c_{nn} \end{pmatrix}$

$$AC^t = \begin{pmatrix} a_{11} & \cdots & a_{1n} \\ \vdots & & \vdots \\ a_{n1} & \cdots & a_{nn} \end{pmatrix} \begin{pmatrix} c_{11} & \cdots & c_{n1} \\ \vdots & & \vdots \\ c_{1n} & \cdots & c_{nn} \end{pmatrix} = \begin{pmatrix} \sum_{j=1}^n a_{1j}c_{1j} & 0 & \cdots & 0 \\ & \ddots & & \vdots \\ & & \ddots & 0 \\ \text{symm.} & & & \sum_{j=1}^n a_{nj}c_{nj} \end{pmatrix} \quad (\because \sum_{j=1}^n a_{ij}c_{kj} = 0, \quad i \neq k)$$

$$= \begin{pmatrix} |A| & 0 & \cdots & 0 \\ & |A| & & \vdots \\ & & \ddots & 0 \\ \text{symm.} & & & |A| \end{pmatrix} = |A| I$$

Lemma

$$\therefore |A| \neq 0 \Rightarrow A \frac{C^t}{|A|} = I \Leftrightarrow A^{-1} = \frac{C^t}{|A|} \quad (\text{of course } C^t A = |A| I \Leftrightarrow \frac{C^t}{|A|} A = I)$$

■ Cramer's rule

$$A\underline{b} = \underline{c}$$

$$|A| \neq 0 \Rightarrow \exists A^{-1} = \frac{1}{|A|} C^t$$

$$\underline{b} = A^{-1} \underline{c} = \frac{1}{|A|} C^t \underline{c} \Rightarrow b_k = \frac{1}{|A|} |A_k|, k = 1, \dots, n$$

$$\therefore \underline{b} = \frac{1}{|A|} \begin{pmatrix} c_{11} & \dots & c_{n1} \\ \vdots & & \vdots \\ c_{1k} & \dots & c_{nk} \\ \vdots & & \vdots \\ c_{1n} & \dots & c_{nn} \end{pmatrix} \begin{pmatrix} c_1 \\ \vdots \\ c_n \end{pmatrix}$$

$$= \frac{1}{|A|} \begin{pmatrix} c_{11}c_1 + \dots + c_{n1}c_n \\ \vdots \\ c_{1k}c_1 + \dots + c_{nk}c_n \\ \vdots \\ c_{1n}c_1 + \dots + c_{nn}c_n \end{pmatrix} = \frac{1}{|A|} \begin{pmatrix} \sum_{j=1}^n c_{j1}c_j \\ \vdots \\ \sum_{j=1}^n c_{jk}c_j \\ \vdots \\ \sum_{j=1}^n c_{jn}c_j \end{pmatrix}$$

$$\therefore b_k = \frac{1}{|A|} \sum_{j=1}^n c_{jk} c_j$$

$$\sum_{j=1}^n c_{jk} c_j = c_{1k} c_1 + \dots + c_{nk} c_n$$

$$= c_1 \times (-1)^{1+k} |M_{1k}| + \dots + c_n \times (-1)^{n+k} |M_{nk}|$$

$$= |A_k|$$

with A_k is a matrix replacing the k^{th} column of A by $\underline{c}$

Note

$$A = \begin{pmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \\ a_{31} & a_{32} & a_{33} \end{pmatrix}, \quad c = \begin{pmatrix} c_1 \\ c_2 \\ c_3 \end{pmatrix}$$

$$\Rightarrow \sum_{j=1}^3 c_{j1} c_j = c_{11} c_1 + c_{21} c_2 + c_{31} c_3$$

$$= c_1 \times (-1)^{1+1} |M_{11}| + c_2 \times (-1)^{2+1} |M_{21}| + c_3 \times (-1)^{3+1} |M_{31}|$$

$$= \begin{vmatrix} c_1 & a_{12} & a_{13} \\ c_2 & a_{22} & a_{23} \\ c_3 & a_{32} & a_{33} \end{vmatrix} = |A_1|$$

$$\Rightarrow b_1 = \frac{1}{|A|} |A_1|$$

■ Properties of $|A|$

$$\text{i) } A = \begin{pmatrix} a_{11} & \cdots & a_{1n} \\ & \ddots & \vdots \\ \underline{0} & & a_{nn} \end{pmatrix} : \text{upper triangular matrix} \quad , \quad A = \begin{pmatrix} a_{11} & & \underline{0} \\ \vdots & \ddots & \\ a_{n1} & \cdots & a_{nn} \end{pmatrix} : \text{lower triangular matrix}$$

$$\Rightarrow |A| = a_{11} \times \cdots \times a_{nn}$$

$$\text{ii) } |AB| = |A||B|$$

$$\text{iii) } |A| = |A^t|$$

$$\text{iv) } \begin{vmatrix} A_{11} & \underline{0} \\ \underline{0} & A_{22} \end{vmatrix} = |A_{11}||A_{22}|$$

$$\text{v) } A_{n \times n} : \text{The orthogonal matrix} \Leftrightarrow A^t A = I_n \text{ with } \underline{a}_j^t \underline{a}_k = 0, \quad j \neq k \text{ and } \underline{a}_j^t \underline{a}_j = 1$$

$$\bullet \quad |A| = \pm 1 \quad (\because 1 = |I_n| = |A^t A| = |A^t| |A| = |A|^2)$$

$$\bullet \quad B_{n \times n} : \text{Any matrix} \Rightarrow |A^t B A| = |B|$$

vi) $|cA| = c^n |A|$ if $A_{n \times n}$ & c is constant

$$\text{vii) } A_{n \times n} = \begin{pmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{pmatrix} \begin{matrix} n_1 \times n_1 & n_1 \times (n-n_1) \\ (n-n_1) \times n_1 & (n-n_1) \times (n-n_1) \end{matrix} : \text{Partitioned matrix}$$

$$|A| = |A_{11}| |A_{22} - A_{21} A_{11}^{-1} A_{12}| \quad \text{if } |A_{11}| \neq 0$$

$$|A| = |A_{22}| |A_{11} - A_{12} A_{22}^{-1} A_{21}| \quad \text{if } |A_{22}| \neq 0$$

viii)* $\lambda_j, j=1, \dots, n$: eigenvalues of $A_{n \times n}$

$$\bullet \quad |A| = \prod_{j=1}^n \lambda_j$$

$$\bullet \quad \text{tr}(A) = \sum_{j=1}^n \lambda_j$$

Application 1

$$\underline{x} = (x_1, x_2, \dots, x_p)^t \sim N_p(\underline{\mu}, \Sigma)$$

$$\underline{x}_{p \times 1} = \begin{pmatrix} \underline{x}_1 \\ \underline{x}_2 \end{pmatrix}_{\substack{q \\ p-q=r}} \sim \underline{\mu} = \begin{pmatrix} \underline{\mu}_1 \\ \underline{\mu}_2 \end{pmatrix}, \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}_{\substack{q \times q & q \times r \\ r \times q & r \times r}}$$

$$\underline{x}_1 \perp \underline{x}_2 \Leftrightarrow \Sigma_{12} = \underline{0}$$

$$\bullet \quad f(\underline{x}) = (2\pi)^{-p/2} |\Sigma|^{-1/2} \exp\left[-\frac{1}{2}(\underline{x} - \underline{\mu})^t \Sigma^{-1}(\underline{x} - \underline{\mu})\right]$$

$$\text{i) } |\Sigma| = |\Sigma_{11}| |\Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}| = |\Sigma_{11}| |\Sigma_{22}| \quad (\because \Sigma_{12} = \underline{0})$$

$$\text{ii) } (\underline{x} - \underline{\mu})^t \Sigma^{-1}(\underline{x} - \underline{\mu}) = \begin{pmatrix} (\underline{x}_1 - \underline{\mu}_1)^t & (\underline{x}_2 - \underline{\mu}_2)^t \end{pmatrix} \begin{pmatrix} \Sigma_{11}^{-1} & \underline{0} \\ \underline{0} & \Sigma_{22}^{-1} \end{pmatrix} \begin{pmatrix} \underline{x}_1 - \underline{\mu}_1 \\ \underline{x}_2 - \underline{\mu}_2 \end{pmatrix}$$

$$= (\underline{x}_1 - \underline{\mu}_1)^t \Sigma_{11}^{-1}(\underline{x}_1 - \underline{\mu}_1) + (\underline{x}_2 - \underline{\mu}_2)^t \Sigma_{22}^{-1}(\underline{x}_2 - \underline{\mu}_2)$$

$$\begin{aligned}
\text{iii) } f(\underline{x}) &= (2\pi)^{-q/2} |\Sigma_{11}|^{-1/2} \exp\left[-\frac{1}{2}(\underline{x}_1 - \underline{\mu}_1)^t \Sigma_{11}^{-1}(\underline{x}_1 - \underline{\mu}_1)\right] \\
&\quad \times (2\pi)^{-r/2} |\Sigma_{22}|^{-1/2} \exp\left[-\frac{1}{2}(\underline{x}_2 - \underline{\mu}_2)^t \Sigma_{22}^{-1}(\underline{x}_2 - \underline{\mu}_2)\right] \\
&= f(\underline{x}_1) \times f(\underline{x}_2)
\end{aligned}$$

$$\underline{x}_1 \sim N_q(\underline{\mu}_1, \Sigma_{11}) \quad \perp \quad \underline{x}_2 \sim N_r(\underline{\mu}_2, \Sigma_{22})$$

Application 2

$$\text{LRT for } \begin{cases} H_0 : \underline{\mu} = \underline{\mu}_0 & \Rightarrow \Lambda^{2/n} = \frac{|\hat{\Sigma}|}{|\hat{\Sigma}_0|} = \frac{|\sum (\underline{x}_i - \bar{\underline{x}})(\underline{x}_i - \bar{\underline{x}})^t|}{|\sum (\underline{x}_i - \underline{\mu}_0)(\underline{x}_i - \underline{\mu}_0)^t|} \\ H_0 : \Sigma = \Sigma_0 & \Rightarrow \Lambda = \left(\frac{e}{n}\right)^{np/2} \text{etr}\left(-\frac{1}{2}\Sigma_0^{-1}A\right) \left|\Sigma_0^{-1}A\right| \quad \text{where } A = \sum (\underline{x}_i - \bar{\underline{x}})(\underline{x}_i - \bar{\underline{x}})^t \end{cases}$$

- Multivariate normal distribution $\sim f(\underline{x}) = (2\pi)^{-p/2} |\Sigma|^{-1/2} \exp\left[-\frac{1}{2}(\underline{x} - \underline{\mu})^t \Sigma^{-1}(\underline{x} - \underline{\mu})\right]$
- Likelihood function : $L(\underline{\mu}, \Sigma) = \prod_{i=1}^n f(\underline{x}_i) = (2\pi)^{-np/2} |\Sigma|^{-n/2} \exp\left(-\frac{1}{2} \sum_{i=1}^n (\underline{x}_i - \underline{\mu})^t \Sigma^{-1}(\underline{x}_i - \underline{\mu})\right)$
- $\Lambda = \max_{\Theta_0} L(\underline{\mu}, \Sigma) / \max_{\Theta_1} L(\underline{\mu}, \Sigma) \Rightarrow -2\ln \Lambda \sim \chi^2 \text{ deg} = \dim(\Theta) - \dim(\Theta_0) \quad \text{where } \Theta = \Theta_0 \cup \Theta_1$